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# Lucky in Life, Unlucky in Love?

## The Effect of Random Income Shocks on Marriage and Divorce

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### ABSTRACT

*Economists have long been interested in the extent to which economic resources affect decisions to marry and divorce. However, this issue has been difficult to address empirically due to a lack of exogenous income shocks. We overcome this problem by exploiting the randomness of the Florida Lottery and comparing recipients of large prizes to those of small prizes. Results indicate that while positive income shocks of \$25,000 to \$50,000 do not cause statistically significant or economically meaningful changes in divorce rates, single women are less likely to marry as a result of the additional income.*

### I. Introduction

Economists and other social scientists have devoted significant effort toward determining the relationship between economic resources and marital decisions. Interest began with work by Becker (1981) on how income affects an individual's incentives to marry or divorce. More recently, researchers and policymakers have tried to determine whether income supports or tax credits for low-income households cause individuals to divorce or fail to marry, which many policymakers

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would regard as an unintended consequence (Bitler, Gelbach, Hoynes, and Zavodny 2004; Cain and Wissoker 1990; Groeneveld, Tuma, and Hanan 1980; Hu 2003).

There are several reasons why positive income shocks could affect marital decisions. For married couples, more generous cash transfers may have a stabilization effect and relax financial constraints and arguments that lead to divorce. Indeed, the income shocks received by individuals in this study—\$25,000 to \$50,000—are sufficiently large to pay off the majority of the unsecured debt (\$49,000) owed by the those on the brink of bankruptcy (Hankins, Hoekstra, and Skiba 2010).

On the other hand, increased resources may enable unhappy couples to incur the costs associated with divorce.<sup>1</sup> Schramm (2006) estimates that legal fees incurred during divorce average \$7,000 while the additional housing cost averages just over \$4,000 per year.<sup>2</sup> The short-term costs associated with divorce may be especially important to the extent couples are liquidity-constrained or behave as such due to risk-aversion. Evidence consistent with widespread liquidity constraints is reported by Laibson, Repetto, and Tobacman (2001), who report that at least 63 percent of U.S. households pay interest on credit card debt every month. In addition, other studies have shown how individuals do not smooth consumption as predicted by the permanent income hypothesis (Shapiro 2005; Stephens 2003). Finally, even the seemingly long-term costs of divorce such as the increased cost of supporting two households are more temporary than one might expect; Kreider (2006) reports that 50 percent of individuals who divorce remarry within five years, and some who do not likely cohabit.

The theoretical effect of positive income shocks on marriage rates is also ambiguous. While simple economic models of marriage predict that an increase in resources will make single individuals more attractive as marriage partners, it also makes single life more attractive (Burstein 2007; Moffitt 1992). In addition, more resources may facilitate marriage due to sociocultural ideals for weddings or married life that require greater assets (Edin 2000).

However, empirically determining whether economic resources affect marriage and divorce has been difficult due to a lack of exogenous pure income shocks. To overcome that identification problem, we exploit random variation in the magnitude of cash prizes won in the Florida Lottery.<sup>3</sup> The crucial identifying assumption is that conditional on winning more than \$600 (the threshold at which the names of the winners are recorded) for the first time, the amount won is random. Tests support this identifying assumption: winners of \$25,000 to \$50,000 come from similar neigh-

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1. This theory has been popularized in the press. In an article entitled “Buy Low, Divorce High” published in the *New York Times* on August 12, 2007, one divorcee and beneficiary of the appreciation in the housing market is quoted as stating “Money is freedom. . . . We made enough money to be able to get divorced and support two households.”

2. Both the legal and housing cost estimates are for married couples with children; we were unable to document any other estimates of the average legal and housing costs associated with divorce. The estimate from Schramm (2006) on housing costs comes from Census reports on average housing costs reported for unmarried individuals. Consequently, it may understate the additional cost in the first year since moving to less expensive housing takes time and since the estimate ignores the actual moving costs.

3. In exploiting exogenous variation in income due to lottery winnings, our identification strategy is similar to that used by Imbens, Rubin, and Sacerdote (2001) and Lindahl (2005) to address the effect of income shocks on labor supply and consumption and on health and mortality, respectively.

borhoods as those who win less than \$1,000, and perhaps more importantly, winners of large sums were no more or less likely to marry or divorce *before* winning the lottery than were (future) winners of small sums.

Results indicate that large income shocks significantly reduce the likelihood that single women marry. Specifically, we find that single women are six percentage points less likely to marry in the three years following the positive income shock, which represents a 40 percent decline. This suggests that additional income may remove some incentive for single women to marry, at least over the short-term.

In contrast, large cash transfers do not affect the marriage rates of men or induce couples to divorce. We find that the divorce rate of married recipients of \$25,000 to \$50,000 in the three years following the income shock is between one-half and one percentage point lower than that of recipients of \$1,000, which is small both in absolute terms and relative to the baseline three-year divorce rate of 8.5 percent. Estimates also allow us to rule out large absolute positive net effects of income shocks on divorce rates: even the upper bound of the 95 percent confidence interval implies that fewer than one in 70 married couples will divorce due to the positive income shock of \$25,000 to \$50,000.

## II. Previous Literature

In assessing the role of economic factors in marital decisions, this study joins a significant existing literature. Perhaps the most closely related research is that which examines the impact of men's or women's earnings on divorce (for example, Hoffman and Duncan 1995; Weiss and Willis 1997).<sup>4</sup> However, by design these estimates pick up the effect of income as well as reduced returns to specialization. Furthermore, as Johnson and Skinner (1986) point out, it is hard to determine whether, for example, women work more hours in response to a higher expected probability of divorce, or whether women are more likely to divorce because they earn higher incomes. Similarly, it is difficult to believe that men in typical data sets who later receive positive income shocks did so randomly rather than because of personal qualities such as loyalty or interpersonal skills. This is problematic given that these characteristics could themselves yield stronger marriages, but are not observed by the econometrician. Finally, while perhaps the most promising variation in income came from the randomization of income guarantees in the Negative Income Tax experiments, methodological problems and the complexity of the experiments have made it difficult to determine what aspect of the experiments, if any, affected marital decision-making (for example, Groenveld et al. 1980; Cain and Wissoker 1990).

There is also a literature that examines the impact of income-changing life events on marital decisions. Bitler et al. (2004) and Hu (2003) found that reducing welfare benefits increases the likelihood of divorce, though given the nature of the policy it is impossible to distinguish the effect of the negative income shock from that of the

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4. For an excellent recent summary of the literature on economic resources and marriage and divorce, see Burstein (2007).

increased work requirement. In addition, while Charles and Stephens (2004) report that job displacement results in higher divorce rates, they argue that it is the information conveyed by the job loss that causes the divorce rather than the loss of income. This highlights the difficulty in determining the effect of income *per se* on marital stability in the absence of truly exogenous shocks to income.

### III. Data and Identification Strategy

To identify the effect of income shocks on marriage and divorce, we exploit the variation in income that is caused by the size of lottery prizes in Florida. Specifically, we obtained data on winners from two games: Florida Lotto and Fantasy 5. Florida Lotto allows players to choose six numbers or have the computer select a number for them, while Fantasy 5 is similar except that there are only five numbers. Both games are parimutuel games in which the amount won is determined by how many numbers the winner matched, the total amount spent on that drawing, and how many players won. For example, while very few Florida Lotto players matched six of six numbers and won the average prize of \$6 million, players who matched five of six numbers won an average of \$4,200, though the prizes from some drawings were much larger.

Similarly, in Fantasy 5, players who matched five of five numbers prior to 2001 won an average of \$20,000, though again the actual amount varied widely depending on how many players won relative to how many played. After 2001, players who matched five of five numbers won an average of \$120,000, while players who matched four of five numbers *when no one matched five numbers* won an average of \$900. Finally, while it is possible for individuals to play up to ten times on each card, an analysis of Fantasy 5 winners revealed that no winners had played the same five numbers multiple times on the same card. As a result, although some people are more likely to enter our data than others (that is, those who play more frequently or who play multiple times on a card), conditional on winning \$600 the amount won is unaffected by the number of plays paid for on a given card. Rather, the identification strategy that we employ largely relies on the assumption that conditional on matching five of six numbers (Florida Lotto) or five of five numbers (Fantasy 5) for the first time, one's underlying propensity to marry or divorce is uncorrelated with how many winners there were relative to the number of players. As discussed later, an important advantage of this research design is that it is testable.

The data for the analysis of divorce include every lottery winner in Miami-Dade and Palm Beach counties from 1988 through 2004, during which there were 73,714 individuals who won up to \$50,000, while the data for the marriage analysis include only Miami-Dade County winners.<sup>5</sup> These winners represent everyone who won at least \$600 by playing either Florida Lotto or Fantasy 5,<sup>6</sup> the minimum amount for which records were kept. For each lottery winner, we observe the individual's name

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5. The data set for the analysis of marriage is smaller because we could not access electronic marriage records for Palm Beach County.

6. The Florida Lottery has other games that we did not use because of insufficient variation in the magnitude of the cash prizes won.

and home zip code, the amount won (adjusted for inflation), the date of the drawing, and the lottery game played. Finally, we note that while we do not observe the marital status of lottery players at the time of winning, this will not be a problem so long as the magnitude of the cash prize is determined randomly. That is, there should be no more unmarried recipients of large cash prizes than there are of small cash prizes.<sup>7</sup> However, as discussed later, the fact that not all lottery winners are married or single does change the interpretation of the marriage and divorce rates observed as well as the estimated impacts of winning \$25,000 to \$50,000.

Data on marriage and divorce come from public records in Miami-Dade and Palm Beach counties in Florida, which is an equitable distribution state with respect to the division of marital assets.<sup>8</sup> These data were linked to the lottery winner data on the basis of first and last name and county of residence. However, before doing so, efforts were made to reduce the number of false positive matches made due to common names. Toward that end, we excluded all names that appeared more than once in the 2007 county phone records. In addition, if lottery records indicated that an individual with a unique name from a given county won more than once, we then use only the first time that individual won.<sup>9</sup> Importantly, this also means that our identifying assumption is that conditional on winning more than \$600 *for the first time*, the amount won is random. We emphasize this because although whether an individual *ever* wins a large prize clearly depends on frequency of play, the magnitude of the *first* prize won does not.

As shown in Tables A1 and A2 in the appendix, eliminating individuals with common names leaves 40,198 lottery winners for the divorce analysis and 26,629 winners for the marriage analysis. As shown in Table A1, restricting the sample in this way does not change the overall distribution of large and small winners, consistent with what one would expect from excluding common names. Moreover, we later test that this sample restriction does not undermine internal validity by showing the similarity in both the neighborhood characteristics and pre-winning marriage and divorce rates of large and small winners.

The remaining individuals reflect a frequency-of-play-weighted population of lottery players in South Florida who have relatively uncommon names, the latter of which likely means our sample underweights Hispanics in South Florida. Still, comparing the neighborhood characteristics of our sample to the United States indicates that our sample is more Hispanic (42 versus 15.4 percent) and somewhat more Black (19 versus 13 percent) than the United States overall.<sup>10</sup> The neighborhood family income of our sample is somewhat lower than that of Miami-Dade and Palm Beach counties (\$40,400 versus \$44,800), which in turn is lower than all of the United

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7. To the extent that income shocks affect marriage or divorce decisions in the short term, this may not be true in the longer term. For example, if positive cash shocks induce married couples to divorce, we may well expect higher marriage rates (per lottery player) for that group in the longer term as individuals remarry. However, since we only examine divorce rates up to three years afterward and find no effect, we are less concerned that our marriage estimates would be biased.

8. This means that in the absence of other considerations (for example, children), a prize received by one spouse should be split equally. Consistent with this law, we find little differences in the effects based on whether the recipient was male or female.

9. Results are unchanged when these individuals are excluded from the analysis.

10. Source: <http://quickfacts.census.gov/qfd/states/12000.html> and authors' calculations.

**Table 1**  
*Lottery Winners Linked to Marriage and Divorce Records after Winning*

Amount Won	Within Three Years of Winning			Within Three Years of Winning		
	Did Not Divorce	Divorced	Percent Divorced	Did Not Marry	Married	Percent Married
< \$1,000	3,478	156	4.29	2,344	218	8.51
\$1,000–\$5,000	20,515	868	4.06	12,907	1,073	7.68
\$5,000–\$10,000	9,268	363	3.77	5,652	468	7.65
\$10,000–\$15,000	1,149	42	3.53	761	54	6.63
\$15,000–\$25,000	1,988	90	4.33	1,399	124	8.14
\$25,000–\$35,000	1,460	64	4.20	1,002	96	8.74
\$35,000–\$50,000	732	25	3.30	500	31	5.84
Total	38,590	1,608	4.00	24,565	2,064	7.75

States (\$50,000).<sup>11</sup> We note, however, that since we do observe individual-level income, we cannot rule out the possibility that the incomes of individuals in our sample are systematically above or below the average for their neighborhood.

Other evidence on the composition of lottery players comes from two sources. Clotfelter, Cook, Edell, and Moore (1999) analyze data from a nationally representative sample of 2,400 adults and report that frequent lottery players are more likely to have lower incomes and less education than less frequent players. Kearney (2005) reports that while frequency of play is approximately equal across the income distribution, college-educated individuals play approximately 40 percent less than high school graduates. In light of these studies and our own data, our judgment is that individuals in our sample are less well educated and likely have lower incomes than the average American. Thus, our findings are most likely to lead to externally valid predictions of the effect of pure income shocks on populations that are less well educated and somewhat poorer than the U.S. average.

The lottery winners in our sample were linked to divorce and marriage records filed in their respective county in the three years prior to and the three years after winning.<sup>12</sup> This was performed via an automated search of each winner's first and last name on the Miami-Dade County and Palm Beach County Clerk of the Court websites.

The results are shown in Table 1. Of the 40,198 winners, 1,608 (4.00 percent) were linked to a divorce case within three years after winning the lottery. For the marriage analysis, 7.75 percent of recipients were linked to a marriage license in the three years after winning.

11. Source: [http://factfinder.census.gov/home/saff/main.html?\\_lang=en](http://factfinder.census.gov/home/saff/main.html?_lang=en) and authors' calculations.

12. Importantly, the indicator used for marriage was the marriage license itself, which is typically obtained no more than one week prior to the wedding itself.

While it is possible that some type I<sup>13</sup> and type II errors were made in linking the lottery winners to divorce and marriage filings, neither type of error should invalidate the research design due to the randomness of amount won. That is, we should be no more or less likely to match winners of large jackpots to divorce or marriage records than we are to match winners of small jackpots, except for the causal effect of the income shock on marital decisions. Consequently, any difference in the marriage or divorce rates of small winners and large winners is properly interpreted as the causal effect of receiving a large cash prize.

One way to check our matching algorithm is to compare the divorce rates implied by our data to those from the Census. To do so, it is important to note that not all of the lottery winners are married. While we know of no measure capturing the proportion of lottery players or winners who are married, Clotfelter et al. (1999) report that lottery participation rates are approximately equal across divorced, married, and unmarried individuals.<sup>14</sup> A back-of-the-envelope calculation based on that result and the demographic characteristics of residents of Miami-Dade and Palm Beach counties suggest that approximately 54 percent of the individuals in our data were married at the time they won the lottery.<sup>15</sup> Consequently, the annual divorce rate as a proportion of married individuals who received less than \$1,000 is (1.43/0.54) 2.65 percent per year. By comparison, 2.35 percent of married Floridians divorced in 2000 according to Census statistics.

An important advantage of our research design is that we can test the identifying assumption that the amount won is uncorrelated with the underlying propensity to marry or divorce, conditional on winning at least \$600 for the first time. First, we show that the amount won is not explained by the winners' neighborhood characteristics. Second, we show that marriage and divorce rates *prior to* winning the lottery are unaffected by the amount (later) won. Third, we show that the estimates are unaffected by the inclusion of zip code fixed effects. Fourth, we show that the estimates are robust to the inclusion of controls capturing the total payout from the current and previous drawings as well as the maximum payout from the previous drawing.

Finally, while our data set offers the distinct advantage of allowing us to identify the causal effect of income shocks on marital decisions, it is important to note the potential limitations of our approach. First, we cannot test whether the marital re-

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13. The degree of Type I error likely depends on the prevalence of misspelled names and nicknames in the data sets. In a check of 200 random names from divorces filed in the county and 200 names of unique lottery winners, we found seven potential nicknames in the lottery data and 17 potential nicknames in the divorce data. These names include names such as Jill, Charlie, Danny, Willy, Betty, Fred, and Steve. While these names would only be a problem if the individual used one name (for example, Jill) when filing for divorce and another (for example, Jillian) when claiming lottery winnings, a conservative estimate is that we should match  $183/200 * 193/200 = 88.3$  percent of individuals who divorce or marry. We found no obvious misspellings among the 400 names, which is not unexpected given the official nature of both data sets.

14. Specifically, they found that the participation rates among married, single, and divorced individuals were 49.7 percent, 52.8 percent, and 56.7 percent, respectively.

15. According to the 2000 Census data for Miami-Dade and Palm Beach counties, 54.27 percent of residents over the age of 18 were married. Assuming equal participation rates by marital group as roughly found by Clotfelter et al. (1999), this means that one could expect that 54.27 percent of lottery winners would be married.

sponse of lottery players is representative of some other population of interest. Second, it is possible that individuals may respond differently to the large cash transfers observed in our data than they would to smaller cash transfers over a longer period of time. Lastly, we cannot address whether income shocks greater than \$50,000 would affect marital decisions, though we note that the size of prizes examined here is nearly one year's income for the average family (\$44,800) and twice the per-capita income of single individuals (\$22,000) from Miami-Dade and Palm Beach counties.<sup>16</sup>

#### IV. Methodology

Given the straightforward nature of this research design, we begin by examining whether the average marriage and divorce rates are different for large winners (those who win \$25,000–\$50,000) than for small winners (those who win between \$600 and \$1,000).<sup>17,18</sup> We focus primarily on marriage and divorce rates in the three years after the individual or couple receives the income shock to maximize the size of the sample, though we note that conclusions are unchanged when we examine divorce and marriage rates in the five years after winning.

Our primary test of whether income shocks cause individuals to marry or divorce is to estimate the following equation using ordinary least squares regression, though similar marginal effect estimates result from using a probit:<sup>19</sup>

$$(1) \quad Outcome_i = \beta_0 + \beta_1(\$1k\text{--}\$10k \text{ Winner})_i + \beta_2(\$10k\text{--}\$25k \text{ Winner})_i \\ + \beta_3(\$25k\text{--}\$50k \text{ Winner})_i + \beta_4(\text{Fantasy } 5)_i \\ + \beta_5(\text{Previous Drawing Controls}) + \Phi_{year} + \theta_{zip \text{ code}} + \varepsilon_i$$

Where  $Outcome_i$  is an indicator variable equal to 1 if lottery player  $i$  divorced (married) within three years after winning the lottery, and  $(\$25k\text{--}\$50k \text{ Winner})$  is an indicator variable equal to one if the individual won between \$25,000 and \$50,000,

16. Source: [http://factfinder.census.gov/home/saff/main.html?\\_lang=en](http://factfinder.census.gov/home/saff/main.html?_lang=en) and authors' calculations.

17. We only examine winners of amounts up to \$50,000 because there were only several hundred unique winners of more than \$50,000, many of whom won between several hundred thousand and several million dollars. This means we are focusing on income shocks likely to be viewed as temporary than permanent. However, results from including all winners of more than \$25,000 are qualitatively similar. For example, results with the full sample indicate that while there is still no effect of income shocks on divorce, receiving more than \$25,000 reduces marriage rates for women by 2.4 percentage points compared to 3.5 percentage points as reported in Table 2, both of which are statistically significant at the 5 percent level.

18. While we would ideally be able to examine the effect of winning even larger amounts than \$25,000–\$50,000, we note that this amount is large relative to the average family income in Miami-Dade and Palm Beach counties of \$44,800. Consequently, it seems reasonable that if higher income causes a net “independence” effect due to relaxed liquidity constraints or the ability to afford two households for a period of time, we would be able to observe that effect in our sample.

19. Probit estimates are available upon request. Although probit or logit can be preferable to ordinary least squares when most values are near zero or one, there are also disadvantages that one might expect to be especially worrisome in this context. For example, omitted variables can cause bias in probit or logit estimations even if they are orthogonal to the treatment variable. In addition, even classical measurement error in the dependent variable can result in inconsistent estimates (Hausman 2001; Hausman et al. 1998).



where the excluded group is those who won between \$600 and \$1,000. The variable (*Fantasy 5*) is an indicator equal to one if the individual won by playing Fantasy 5 (where Florida Lotto is the excluded game). This allows for a level shift that accounts for any time-invariant difference in the underlying propensity to marry or divorce between Fantasy 5 and Lotto players.<sup>20</sup>  $\Phi$  is a vector of year fixed effects, while  $\theta$  is a vector of zip code fixed effects. The variable (*Previous Drawing Controls*) is a vector of controls for the total payout from the previous drawing and the maximum prize won in the previous drawing, which control for any potential selection into playing based on past pot or prize size.

The primary coefficient of interest is  $\beta_3$ , the sign of which indicates whether receiving large positive income shocks induces individuals to be more or less likely to marry or divorce.

## V. Results

### A. Tests of the Identification Strategy

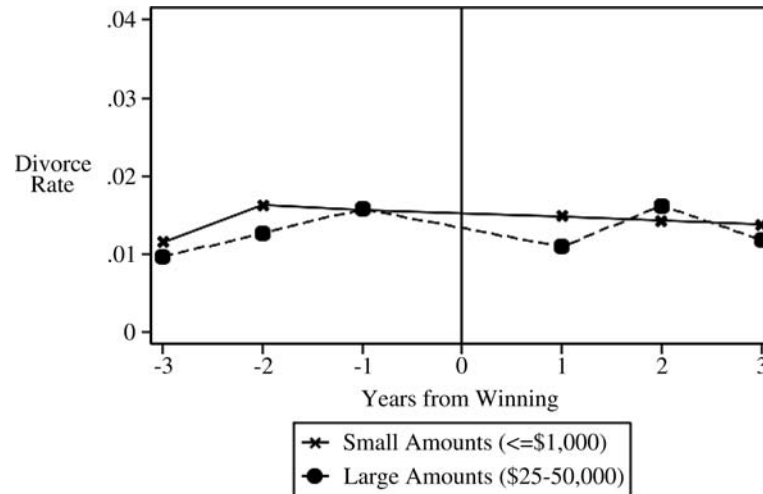
To demonstrate that the size of the income shock received is uncorrelated with other determinants of marriage and divorce, we offer two tests. First, in results available upon request, we regress the cash prize on 13 neighborhood demographic characteristics measuring income, gender, marital status, and educational attainment (as well as interactions) and find that only one coefficient is statistically significant at the 5 percent level.<sup>21</sup> More importantly, all of the variables collectively explain only 0.1 percent of the variation in lottery winnings for individuals in Miami-Dade and Palm Beach counties.

Second, we match lottery winners to marriage and divorce records filed in the years *before* the individual won the lottery. In the absence of correlation between amount won and underlying propensity to marry or divorce, we would expect no difference between the marriage and divorce rates of individuals who would *later* win a large cash prize compared to those of individuals who would *later* win a small cash prize. The estimated impacts of winning \$25,000 to \$50,000 relative to winning less than \$1,000 are reported in Table A3 in the web appendix, and include estimates for the full sample as well as separate estimates for men and women. Of the 24 estimates reported, none are statistically significant at the 5 percent level, and only three are significant at the 10 percent level, which is approximately what one would expect by chance.<sup>22</sup>

20. We also estimated the model using only Fantasy 5, which has both large and small winners. Estimates of the effect of large income shocks on women's marriage rates range from statistically significant  $-2.6$  to  $-4.4$  percentage points, compared to approximately  $-3.5$  percentage points reported in Table 3.

21. The significant variable was the proportion of women from the county, the coefficient of which implies that a five-percentage-point increase in the proportion of women is associated with a prize that is \$325 larger. We note that this is small relative to the amounts examined in this paper.

22. These are three of the four estimates of the effect of (later) winning a large prize on women's divorce rates.



**Figure 1**  
*Flows into Divorce before and after Receiving Small and Large Cash Prizes*

### *B. The Effect of Positive Income Shocks on Divorce*

We now turn to the effect of income shocks on divorce. We begin by examining whether flows into divorce for small winners (those who won less than \$1,000) are different from large winners (those who won between \$25,000 and \$50,000) after winning. Results are shown in Figure 1, which shows that there is little difference between the divorce rates of these two groups before or after the income shocks.

To formally test for differences in the divorce rates of the two groups in the three years after winning, we estimate Equation 1 from above. Results are shown in Table 2, in which Specification 1 indicates that the raw difference between winning \$25,000 to \$50,000 and winning less than \$1,000 is  $-0.39$  percentage points, which is relative to a baseline divorce rate of 4.3 percent. Estimates of the effect of winning \$1,000 to \$10,000 and \$10,000 to \$25,000 are similarly small and statistically insignificant. Including controls for game, year, neighborhood, and current and previous drawing payouts barely changes the estimate, as shown in Specifications 2 through 4. None of the estimates are statistically significant.

Because not all lottery players were married at the time of the income shock, the estimates must be reweighted in order to interpret the estimates as divorce rates relative to the married population. According to data from a nationally representative survey reported by Clotfelter et al. (1999) and Miami-Dade and Palm Beach county demographics, we estimate that 54 percent of lottery winners were married at the time they won. Consequently, the 4.3 percent baseline divorce rate in the lottery winner population over the three years after winning corresponds to a  $(4.3/0.54)$  7.96 percent actual divorce rate among married individuals. Similarly, the  $-0.39$  percentage point difference implies a statistically insignificant 0.72 percentage point

**Table 2**  
*The Effect of Receiving Up to \$50,000 on Marriage and Divorce (relative to receiving less than \$1,000)*

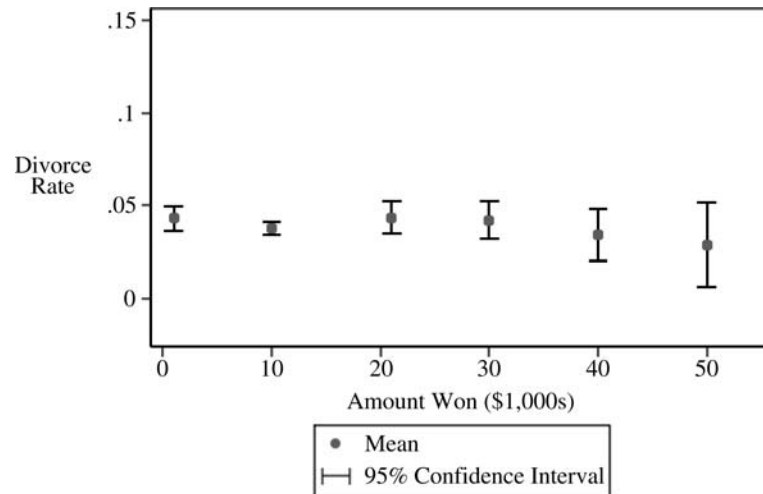
Specification	(1)	(2)	(3)	(4)
<b>Panel 1: The Effect of Random Income Shocks on Divorce Rates within Three Years</b>				
\$1,000–\$10,000	–0.0032 (0.0035)	0.0024 (0.0040)	0.0021 (0.0032)	0.0010 (0.0033)
\$10,000–\$25,000	–0.0025 (0.0048)	–0.0029 (0.0051)	–0.0027 (0.0043)	–0.0037 (0.0044)
\$25,000–\$50,000	–0.0039 (0.0053)	–0.0038 (0.0058)	–0.0032 (0.0056)	–0.0039 (0.0057)
Observations	40,198	40,198	40,198	38,383
<b>Panel 2: The Effect of Random Income Shocks on Marriage Rates within Three Years</b>				
\$1,000–\$10,000	–0.0084 (0.0058)	–0.0060 (0.0064)	–0.0060 (0.0065)	–0.0072 (0.0065)
\$10,000–\$25,000	–0.0090 (0.0078)	–0.0093 (0.0083)	–0.0089 (0.0079)	–0.0099 (0.0085)
\$25,000–\$50,000	–0.0071 (0.0086)	–0.0113 (0.0095)	–0.0107 (0.0095)	–0.0126 (0.0097)
Observations	26,629	26,629	26,629	25,539
Includes game and year fixed effects?	No	Yes	Yes	Yes
Includes neighborhood fixed effects?	No	No	Yes	Yes
Controls for maximum prize and total payout from previous drawing	No	No	No	Yes

Notes: Robust standard errors are in parentheses. Estimates are relative to receiving less than \$1,000 and compare to three-year divorce and marriage rates of 4.3 percent and 8.5 percent, respectively.

\* Significant at the 10 percent level

\*\* Significant at the 5 percent level

\*\*\* Significant at the 1 percent level



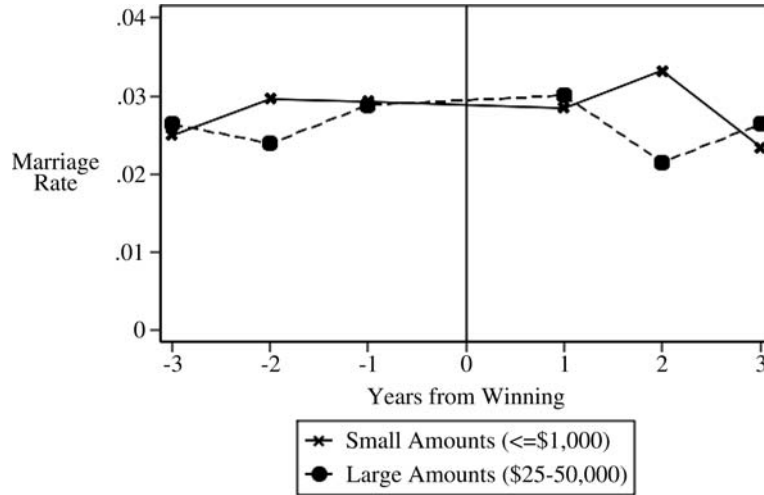
**Figure 2**  
*Observed Divorce Rates in the 3 Years after Winning the Lottery*

drop in the divorce rate due to winning \$25,000–\$50,000. Importantly, while the estimates are not statistically different from zero, they are reasonably precise. The upper bound of the 95 percent confidence interval suggests that no more than one in 70 married couples will be induced to divorce as a result of receiving a large income shock.<sup>23</sup> This implies that to the extent life events or policy changes affect divorce, it is unlikely due to a pure positive income shock.

To ensure that the results are not driven by the admittedly arbitrary cutoffs made in defining large and small income shocks, divorce rates for the full range of income shocks are shown in Figure 2. Consistent with the estimates in Table 2, there appears to be no relationship between the magnitude of the income shock and the likelihood of subsequently filing for divorce.

In results available upon request, we also find no difference between winners coming from minority- or majority-Hispanic zip codes, or above or below median income zip codes. We also tested whether the effect of income shocks on divorce was different by gender, which we inferred on the basis of first names using empir-

23. The reweighted point estimate is  $-0.0072$ , while the reweighted standard error is  $0.0057/0.54 = 0.0106$ . Consequently, the upper bound for the 95 percent confidence interval is  $-0.0072 + 1.96 \cdot 0.0106 = 0.0136$ . In absolute terms, this implies that fewer than one in 70 married couples will divorce as a result of receiving the income shock, while in relative terms it implies that there will be no more than a 16 percent ( $0.0136/0.086$ ) increase in the divorce rate among married lottery players.



**Figure 3**  
*Flows into Marriage before and after Receiving Small and Large Cash Prizes*

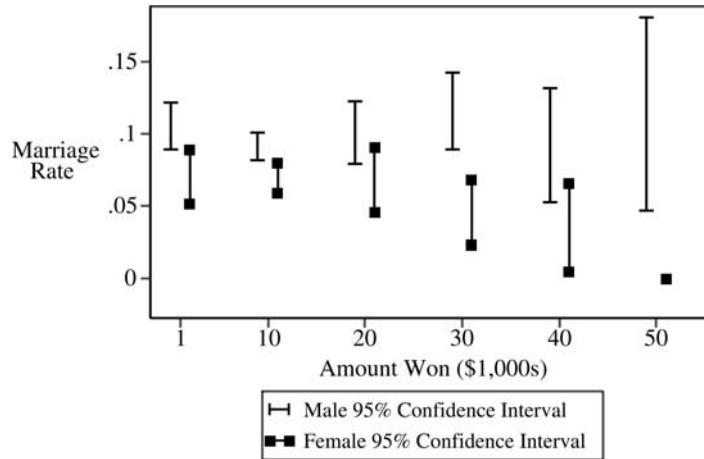
ical distributions of Florida schoolchildren generously provided by David Figlio.<sup>24</sup> Results (available upon request) indicate that neither men nor women respond to income shocks on the divorce margin. The similarity of the (non)responses is consistent with the Coase Theorem, which predicts that changing the initial allocation of resources will only affect the settlements received by each partner, rather than the final outcome. Finally, we note that using probit to estimate the model rather than OLS yielded similar estimates.<sup>25</sup>

**C. The Effect of Positive Income Shocks on Marriage**

We now turn to the question of whether large positive income shocks affect an individual’s likelihood of getting married. Flows into marriage are shown in Figure 3, which indicates that consistent with the identifying assumption, there is little

24. Using these empirical probability distributions, we were able to categorize the gender of approximately 83 and 81 percent of winners in the divorce and marriage data sets, respectively. Of those names we could not identify, most appeared to be foreign in origin and thus did not appear in the data set of schoolchildren’s first names provided by David Figlio. When we estimated the effect of large income shocks on the marriage and divorce rates of this group, we found no effect on marriage and a marginally significant effect on divorce ranging from -1 to -1.6 percentage points. The marriage and divorce rates of the unidentified individuals are 40 and 15 percent lower than for those whose gender we can identify, respectively.

25. For example, results from a probit estimation that include game and year fixed effects imply that winning \$25,000 to \$50,000 causes divorce rates over the next three years to decrease by a statistically insignificant 0.22 percentage points, compared to a reduction of 0.32 percentage points from the OLS regression. Similarly, the effect on women’s marriage rates in the three years after winning is -2.5 percentage points compared to an OLS estimate of -3.2 percentage points, both of which are significant at the 5 percent level.



**Figure 4**  
*Observed Marriage Rates for Men and Women in the 3 Years after Winning the Lottery*

difference in the marriage rates prior to the income shock. After the \$25,000 to \$50,000 income shock, Figure 3 also shows no definitive evidence of effects on marriage rates in the 3 years afterward. Results from more formal tests of the effect of winning \$25,000 to \$50,000 on marriage rates within 3 years are shown in Table 2. The marriage rates of winners of \$25,000 to \$50,000 are 0.7 percentage points lower than those who received less than \$1,000. Adding controls increases the magnitude of the estimate to 1.3 percentage points. All estimates are relative to baseline marriage rates among lottery players of 8.5 percent.

To further investigate this result, we examine whether women respond differently than men.<sup>26</sup> To do so, we used empirical distributions on the gender of first names of Florida schoolchildren provided by David Figlio to categorize 21,585 of the 26,629 individuals in the marriage data set as either male or female. The results are shown graphically in Figure 4,<sup>27</sup> while the regression estimates are shown in Table 3.<sup>28</sup> Both indicate that while positive income shocks do not affect men's marriage rates, women who receive \$25,000 to \$50,000 are significantly less likely to marry than women who receive less than \$1,000. Point estimates in Specifications 1

26. We also investigate differences by income and race as proxied by the demographics of each winner's zip code and found no differences.

27. The reader will note that male lottery players appear more likely to marry than female players. We emphasize, however, that this may be because male players are more likely to be single than female players.

28. In addition to the semiparametric approach used in Table 3, we also regressed women's marriage outcomes on a cubic function of amount won. Results are similar to those shown visually in Figure 4 and formally in Table 3. Winning less than \$30,000 did not have a statistically significant impact on marriage, while the predicted impact of winning \$30,000 and \$40,000 relative to \$1,000 were  $-1.9$  and  $-3.7$  percentage points, respectively, both of which are statistically significant at the 5 percent level.

**Table 3**  
*The Effect of Receiving \$25,000 to \$50,000 on Three-Year Marriage Rates for Men and Women*

Specification	(1)	(2)	(3)	(4)
Panel 1: Women				
\$1,000–\$10,000	–0.0024 (0.0100)	–0.0011 (0.0105)	0.0010 (0.0104)	–0.0005 (0.0106)
\$10,000–\$25,000	–0.0079 (0.0130)	–0.0092 (0.0135)	–0.0070 (0.0136)	–0.0084 (0.0140)
\$25,000–\$50,000	–0.0293** (0.0130)	–0.0341** (0.0137)	–0.0317** (0.0138)	–0.0342** (0.0140)
Panel 2: Men				
\$1,000–\$10,000	–0.0121 (0.0089)	–0.0112 (0.0095)	–0.0118 (0.0095)	–0.0126 (0.0096)
\$10,000–\$25,000	–0.0123 (0.0119)	–0.0135 (0.0124)	–0.0139 (0.0124)	–0.0141 (0.0128)
\$25,000–\$50,000	0.0026 (0.0138)	–0.0020 (0.0145)	–0.0007 (0.0146)	–0.0024 (0.0147)
Observations	21,585	21,585	21,585	20,678
Includes game and year fixed effects?	No	Yes	Yes	Yes
Includes neighborhood fixed effects?	No	No	Yes	Yes
Controls for maximum prize and total payout from previous drawing	No	No	No	Yes

Notes: Robust standard errors are in parentheses. Estimates are relative to winning less than \$1,000 and compare to three-year marriage rates of 7.0 percent for women and 10.6 percent for men.

\* Significant at the 10 percent level

\*\* Significant at the 5 percent level

\*\*\* Significant at the 1 percent level

through 4 of Table 3 indicate that women who receive the positive income shock are between 2.9 and 3.4 percentage points less likely to marry in the next three years than their counterparts who received less than \$1,000.<sup>29</sup> These effects are large, representing 41 to 48 percent reductions relative to the baseline marriage rate among all female lottery players of 7 percent.

Additionally, we note that the difference between winning \$25,000 to \$50,000 is statistically distinguishable from the effect of winning \$1,000 to \$10,000 at the 1 percent level in all four specifications of Table 3. Similarly, the difference between winning \$25,000 to \$50,000 and winning \$10,000 to \$25,000 is statistically significant at the 5 percent level for Specifications 2 and 4, and at the 10 percent level for Specifications 1 and 3. This suggests it takes a relatively large income shock to induce single women not to marry.

#### *D. Migration Out of the County of Residence*

The identification strategy utilized in this paper will break down if large income shocks cause individuals to move to another county before marrying or divorcing.<sup>30,31</sup> We expect this to be unlikely for several reasons. First, residents of Miami-Dade and Palm Beach counties appear to have strong roots: the Census reports that over 80 percent of residents in these counties lived in the same county five years earlier, a number which would likely be higher if not for substantial migration into the area over this time period. Second, both Miami-Dade and Palm Beach counties are very large and offer a diverse set of areas in which one could relocate without leaving the county. Geographically, Miami-Dade and Palm Beach counties cover 1,946 and 1,974 square miles, respectively, making each of them over six times the size of New York City and over one-third as large as the state of Connecticut. Both are also large in terms of population; Miami-Dade County was home to 2.3 million people in 2000 while 1.1 million people resided in Palm Beach County. Furthermore, there is substantial within-county variation in neighborhoods in both counties. For example, the median family income in the poorest zip code in Miami-Dade County was \$18,000 while the median family income in the wealthiest zip code was \$200,000. Collectively, these factors minimize the likelihood that a wealth shock will cause individuals to move out of the county.

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29. Graphs showing flows into marriage over time for men and women separately are available in the web appendix. As shown there, the result for women appears to be driven largely by fewer marriages in the second and, to a lesser extent, third year after winning the lottery. While this pattern is not statistically precise, it is suggestive that while additional income may not cause women to cancel weddings that are already planned, it may induce women to either exit or lengthen the duration of relationships less close to marriage.

30. The potential for bias is smaller here than it might otherwise be due to how the sample was constructed. Specifically, because all individuals are assumed to not divorce (marry) unless matched to a divorce (marriage) record, the results are unaffected if large income shocks cause differential attrition only among individuals who will not later divorce (marry).

31. A similar problem might arise if large income shocks induce women to obtain a marriage license in a county other than the one in which they reside (for example, the couple elopes to Las Vegas). While we cannot directly test for this behavior, it seems unlikely to drive our results. That is because while one can certainly afford trips to Las Vegas or most any other place with, say, \$10,000, as shown in Figure 4 we do not observe an effect on marriage for recipients of less than \$30,000.



**Table 4**  
*The Proportion of Lottery Winners Who Are Linked to County Phone Records One to Five Years Later*

	Winners of Less than \$1,000 (1)	Winners of More than \$25,000 (2)	Difference (3) = (2) - (1)
Panel 1: All Winners 0 to 5 Years Later	0.543 (0.007)	0.552 (0.023)	0.009 (0.024)
Panel 2: Men Only 0 to 5 Years Later	0.564 (0.010)	0.565 (0.033)	0.001 (0.034)
Panel 3: Women Only 0 to 5 Years Later	0.536 (0.013)	0.565 (0.042)	0.029 (0.044)

Notes: Robust standard errors are in parentheses.  
None of the differences are statistically different at conventional levels.

We can offer one test of whether differential migration out of the county is likely to be a problem for our analysis. Specifically, we link individuals who won the lottery between March of 2003 and March of 2007 to phone records accessed in March of 2008.<sup>32</sup> Results are reported in Table 4, which shows that the difference in the proportion of small and big winners showing up in the county phone records zero to five years later is small and statistically insignificant. Moreover, back-of-the-envelope calculations indicate that even if we use the lower bound of the 95 percent confidence interval from attrition figures in Table 4, fully 75 percent of attriters from the county would have had to marry in other counties in order to explain the observed drop in marriage rates for women, a marriage rate that is 10 times that of the baseline group.<sup>33</sup> Consequently, while this is an imperfect test due to the fact that some households no longer have landlines, some individuals who do are not listed in the phone book, and winning the lottery could potentially enable some

32. Note that these individuals were not used in the main analysis. We are unable to use this test on the main sample of lottery winners because we do not have access to historical phone records.

33. The smallest estimates from Table 3 indicate a 2.9 percentage point decline in marriage rates within three years, while the lower bound of the 95 percent confidence interval from Table 4 suggests that large winners should be no more than 3.8 percentage points ( $0.9 - 1.96 * 2.4$ ) more likely to leave the county than small winners in the five years after winning. Thus, for the treatment effect estimate to be entirely due to attrition,  $2.9 / 3.8 = 76$  percent of attriters had to have married outside the county. Alternatively, using the attrition figures for women in Table 4 implies that at least 51 percent of attriters had to have married outside the county to explain the result. Either scenario is especially unlikely since only approximately half of winners were unmarried at the time of winning.

families to afford a landline, it does provide comfort that differential migration from the county is unlikely to explain the results.<sup>34</sup>

## VI. Discussion and Conclusions

While economists have long been interested in the relationship between economic resources and marital decisions, determining the extent to which economic resources affect marriage and divorce has been difficult due to a lack of exogenous income shocks. To overcome that problem, we exploit income shocks that occur as a result of winning the lottery and compare the marriage and divorce rates of individuals who won between \$25,000 and \$50,000 to those who won just over \$600. We find no evidence that pure income shocks cause statistically significant or economically meaningful changes in divorce rates. Moreover, even the upper bound of the 95 percent confidence interval implies that fewer than one in 70 married couples will be induced to divorce as a result of receiving an income shock that is approximately twice as large as per-capita income. However, we do find that positive income shocks reduce the likelihood that single women will marry in the following three years by 40 percent.

The combination of a large impact on women's marriage rates and the lack of an effect on divorce rates presents an interesting puzzle. Why, after receiving a large income shock, would women avoid getting married but not exit existing marriages? Indeed, the Coase Theorem suggests that if the income shock causes total gains from being single to exceed total gains from being married, couples should not be married. This apparent contradiction can be resolved in two ways. The first relates to the role of transaction costs in marital decision-making. Although it may be optimal, after receiving the income shock, for married women to have delayed or avoided marriage, significant transaction costs are associated with divorce. In contrast, few, if any, transaction costs are associated with not getting married in the first place.

Alternatively, it may be that upon receiving the income shock, total welfare for a couple is *not* higher unmarried than married, even though the utility of the woman may be higher if she remains single. That is, women may remain single even if that lowers the welfare of others, so long as those other parties (for example, men, children) cannot credibly promise to transfer gains from the marriage to the woman. Indeed, the social welfare implications of our finding depend crucially on why women are marrying at lower rates. On the one hand, large positive income shocks may reduce the risk-sharing benefits of marriage, or make single life more attractive in other ways. In this case, while the additional income would presumably lead to higher utility for the women, it is difficult to know how it would affect the welfare of others, such as prospective partners or children. Alternatively, it is possible that the additional wealth may enable women to postpone marriage in search of a better

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34. These factors are likely the reason why the number of individuals linked to phone records after winning is lower than one would otherwise expect. This is reflected by the fact only 53 percent of winners of less than \$1,000 are found in the phone records in the six months after winning.

match.<sup>35</sup> In that case, the income shocks may be welfare-enhancing for prospective mates and children, as well as the single women themselves.

In examining the impact of exogenous shocks on marriage decisions, our paper fits into a much larger literature that examines the sensitivity of marriage and divorce to economic forces. Results here are broadly consistent with existing research. For example, Charles and Stephens (2004) report that while spousal job loss appears to cause divorce, spousal disability that yields a similar loss in household income does not. Charles and Stephens (2004) interpret this finding as casting doubt on a purely pecuniary motivation for divorce following a negative income shock, an interpretation consistent with the lack of evidence reported in this study. Similarly, while the lack of a finding on divorce in this paper contrasts with studies of welfare reform in which lower income supports are associated with lower divorce rates (Bitler et al. 2004; Hu 2003), evidence presented here suggests that the reduced divorce rates were caused by the rule changes accompanying welfare reform, rather than the change in income per se.

Our finding that random positive income shocks reduce women's marriage rates is also broadly consistent with evidence that marriage decisions are affected by exogenous factors, including the gender of the first-born child (Dahl and Moretti 2008), minimum age requirements (Blank, Charles, and Sallee 2009), and blood test requirements (Buckles, Guldi, and Price 2009). However, our finding contrasts with the findings of other credible research designs that exploit income variation from welfare reform and report either no or weak effects on marriage (Bitler et al. 2004; Hu 2003), or exploit variation from antipoverty experiments and find higher income is correlated with higher marriage rates for single women (Gassman-Pines and Yoshikawa 2006). Similarly, our findings stands in contrast to Burstein's (2007) characterization that studies using microdata usually find positive or no relationships between women's income and marriage (for example, McLaughlin and Lichter 1997; Smock and Manning 1997).

Our finding does not necessarily contradict these studies, however. By design, antipoverty programs and welfare reforms simultaneously changed other determinants of marriage, while earned income observed in other studies is likely correlated with other factors that make women attractive marriage partners. Thus, our interpretation is that while additional income reduces marriage rates for single women *ceteris paribus*, this effect may be offset by factors such as increased work requirements that make the women more attractive mates and/or introduce them to higher-quality single men.

Collectively, our results yield two important conclusions for policy. First, it appears that changing only the income of married couples will be unlikely to have an impact on the divorce margin. This indicates that to the extent policymakers wish to impact divorce rates, they ought to concern themselves primarily with the non-pecuniary determinants of marital instability. In addition, our results indicate policies that provide additional income will tend to reduce the short-term marriage rates of

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35. In other (unreported) results, we found little difference in the marriage rates of large and small female winners in the fourth and fifth years after winning, leaving the size and statistical significance of the point estimates unchanged. Unfortunately, we do not observe enough winners long enough for us to examine marriage rates more than five years after winning.

single women, unless the policy is accompanied by other restrictions such as increased work requirements.

**Table A1**

*Constructing the Sample of Unique Lottery Winners in Miami-Dade and Palm Beach Counties Used for Divorce Analysis*

Amount Won	All Winners		First Time Winners		Unique in Phonebook	
	Number	Percent	Number	Percent	Number	Percent
<\$1,000	7,713	10.46	5,307	9.11	3,634	9.04
\$1,000–\$5,000	39,151	53.11	30,794	52.89	21,383	53.19
\$5,000–\$10,000	16,447	22.31	13,910	23.89	9,631	23.96
\$10,000–\$15,000	2,130	2.89	1,810	3.11	1,191	2.96
\$15,000–\$25,000	3,888	5.27	3,027	5.20	2,078	5.17
\$25,000–\$35,000	2,938	3.99	2,263	3.89	1,524	3.79
\$35,000–\$50,000	1,447	1.96	1,115	1.91	757	1.88
Total	73,714	100.00	58,226	100.00	40,198	100.00

**Table A2**

*Constructing the Sample of Unique Lottery Winners in Miami-Dade County Used for Marriage Analysis*

Amount Won	All Winners		First Time Winners		Unique in Phonebook	
	Number	Percent	Number	Percent	Number	Percent
<\$1,000	5,872	11.04	3,887	9.66	2,562	9.62
\$1,000–\$5,000	27,997	52.62	21,020	52.27	13,980	52.50
\$5,000–\$10,000	11,351	21.33	9,227	22.94	6,120	22.98
\$10,000–\$15,000	1,568	2.95	1,293	3.21	815	3.06
\$15,000–\$25,000	3,034	5.70	2,296	5.71	1,523	5.72
\$25,000–\$35,000	2,294	4.31	1,693	4.21	1,098	4.12
\$35,000–\$50,000	1,090	2.05	802	1.99	531	1.99
Total	53,206	100.00	40,218	100.00	26,629	100.00

**Table A3**  
*Falsification Test: The Impact of (Later) Receiving \$25,000 to \$50,000 on Marriage and Divorce Rates Prior to Winning*

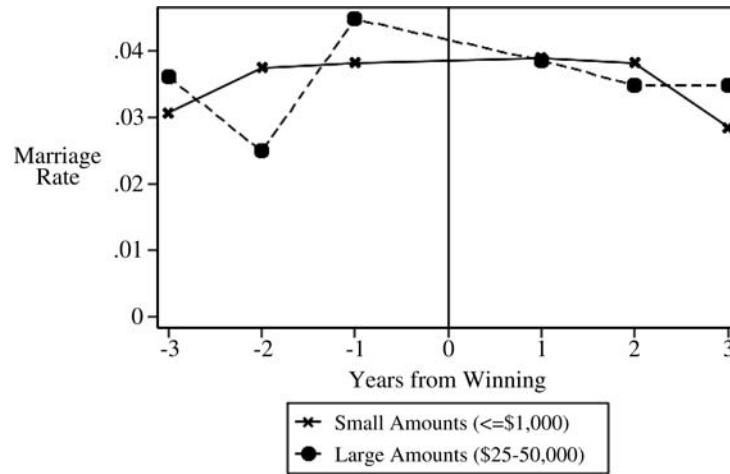
Specification	(1)	(2)	(3)	(4)
<b>Panel 1: Divorce Rates Prior to Winning</b>				
Sample/Subgroup				
All winners	-0.0051 (0.0052)	-0.0053 (0.0057)	-0.0049 (0.0053)	-0.0050 (0.0055)
Men only	0.0012 (0.0082)	0.0009 (0.0086)	0.0016 (0.0085)	0.0015 (0.0086)
Women only	-0.0153* (0.0086)	-0.0160* (0.0089)	-0.0148* (0.0089)	-0.0141 (0.0090)
<b>Panel 2: Marriage Rates Prior to Winning</b>				
Sample/Subgroup				
All winners	-0.0053 (0.0086)	-0.0110 (0.0096)	-0.0105 (0.0078)	-0.0096 (0.0080)
Men only	-0.0019 (0.0137)	-0.0041 (0.0145)	-0.0051 (0.0146)	-0.0056 (0.0147)
Women only	-0.0051 (0.0146)	-0.0077 (0.0153)	-0.0049 (0.0153)	-0.0055 (0.0155)
Includes game and year fixed effects?	No	Yes	Yes	Yes
Includes neighborhood fixed effects?	No	No	Yes	Yes
Controls for maximum prize and total payout from previous drawing	No	No	No	Yes

Notes: Divorce estimates for the full sample includes 40,198 observations; estimates for men and women come from a separate regression on a gender-identified sample of 33,240 observations. Marriage estimates for the full sample come from 26,629 observations; estimates for men and women come from a separate regression on a gender-identified sample of 21,585 observations. Robust standard errors are in parentheses. Estimates are relative to receiving less than \$1,000 and compare to overall three-year divorce and marriage rates of 4.3 percent and 8.5 percent, respectively.

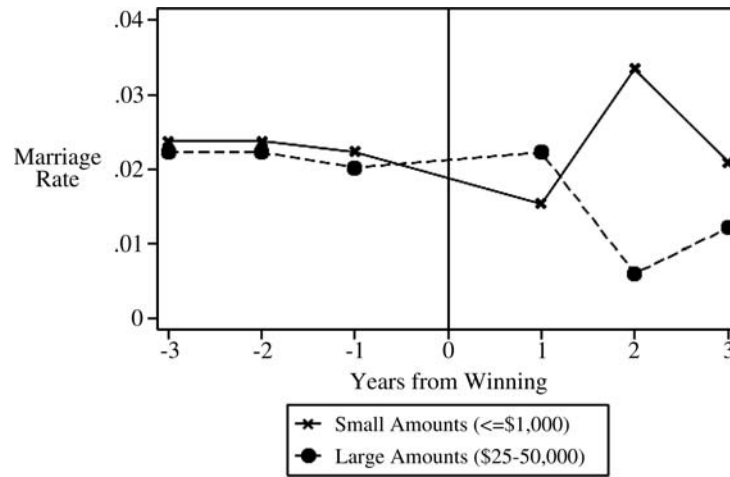
\* Significant at the 10 percent level

\*\* Significant at the 5 percent level

\*\*\* Significant at the 1 percent level



**Figure A1**  
Flows into Marriage for Men before and after Winning the Lottery



**Figure A2**  
Flows into Marriage for Women before and after Winning the Lottery

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